Enumerating Optimal Cost-Constrained Adjustment Sets

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Measuring the Causal Effect



From $P(Y \mid do(a))$ we can compute quantities of interest: ATE: $\mathbb{E}_P(Y \mid do(a)) - \mathbb{E}_P(Y \mid do(a'))$

¹Illustration based on Rotnitzky 2021.

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Can the causal effect of a on y be identified from the observational dist. ?

Yes.
$$P(y \mid do(a)) = \sum_{r} P(y \mid a, r) \cdot P(r)$$

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Measuring the Causal Effect

Observational Dist. + Causal Graph \Rightarrow Interventional Dist. $P(q_2, h_1, r, h_2, a, y)$ $P(Y \mid do(a))$ $P(Y \mid do(a))$

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$$P(y \mid do(a)) = \sum_{r} P(y \mid a, r) \cdot P(r)$$

In this example, $\{r\}$ is an **adjustment set** for estimating the causal effect of *a* on *y*. Since *r* is observed, then $\{r\}$ is a **valid adjustment set**, and hence the causal effect of *a* on *y* is **identifiable**.

¹Illustration based on Rotnitzky 2021.

Adjustment Sets

Definition (Adjustment Sets)

Given a DAG G, and pairwise disjoint X, Y, $Z \subseteq V(G)$, Z is an *adjustment set* for estimating the causal effect of X on Y if, for every distribution P that factorizes according to G:

$$P(\mathbf{Y} \mid do(\mathbf{X})) = \begin{cases} P(\mathbf{Y} \mid \mathbf{X}) & \text{if } Z = \emptyset\\ \sum_{\mathbf{Z}} P(\mathbf{Y} \mid \mathbf{X}, \mathbf{Z}) P(\mathbf{Z}) & \text{otherwise} \end{cases}$$

The choice of adjustment set Z affects:

- 1. Time to compute interventional dist. $P(\mathbf{Y} \mid do(\mathbf{X}))$ and the derived quantities of interest. Takes time exponential in $|\mathbf{Z}|$.
- 2. The variance $\sigma_{Z}^{2}(P)$ of the (asymptotically) normally-distributed estimator $\hat{\mathbb{E}}_{P}(Y \mid do(X))$ Smucler et al. 2021, Smucler and Rotnitzky 2022, Runge 2021.

Natural Requirements from Adjustment Sets

Let G be a DAG with observable variables $R \subseteq V(G)$. We wish to find adjustment sets $Z \subseteq V(G)$ that are:

- 1. Valid : $Z \subseteq R$.
- 2. Non-Redundant (Minimal) : there is no $Z' \subsetneq Z$ that is an adjustment set for estimating the causal effect.
- 3. Small size (or weight); of size at most k.
- 4. Yields estimator with low variance $\sigma_{\mathbf{Z}}^2(P)$:

$$\hat{\mathbb{E}}_{P}(\boldsymbol{Y} \mid do(\boldsymbol{X})) = \sum_{\boldsymbol{Z}} \hat{\mathbb{E}}_{P}(\boldsymbol{Y} \mid do(\boldsymbol{X}), \boldsymbol{Z}) \cdot P(\boldsymbol{Z}).$$

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Examples of Adjustment Sets



Minimal adjustment sets for estimating the causal effect of a on y: $\{t\}$, $\{r\}$, $\{h_1, q_2\}$, $\{h_2, q_1\}$, $\{h_1, h_2\}$, $\{q_1, q_2\}$.

- ► Adjustment set(s) that optimize for size: {*t*}, {*r*}.
- Adjustment set(s) that optimize for variance: {h₁, h₂} Smucler et al. 2021.

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Problem: Size and variance are at odds!

²Illustration based on Rotnitzky 2021.

Domination Among Adjustment Sets

- Let Z₁, Z₂ ⊆ V(G) be two valid adjustment sets for estimating the causal effect of X on Y in G.
- ► We write $Z_1 \leq_G^{\sigma} Z_2$ if $\sigma_{Z_1}^2(P) \leq \sigma_{Z_2}^2(P)$ for every joint probability dist. *P* that factorizes according to *G*.
 - ▶ Induces a partial order. There may exist two adjustment sets S_1 , S_2 where $\sigma_{S_1}^2(P) < \sigma_{S_2}^2(P)$ and $\sigma_{S_2}^2(P') < \sigma_{S_1}^2(P')$ where P, P' both factorize according to G.

• We say that Z_1 dominates Z_2 if:

$$Z_1 \leq_{\mathcal{G}}^{\sigma} Z_2$$
 and $|Z_1| \leq |Z_2|$

and one of these inequalities is proper.

Pareto-Optimal Frontier of Adjustment Sets

Goal: Return all non-dominated adjustment sets.

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Problems:

- 1. Too many.
- 2. Too big/expensive to be practical.

Refined Goal: Return all <u>non-dominated</u> adjustment sets whose size (weight) is at most *k*.

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Refined Goal: Return all <u>non-dominated</u> adjustment sets whose size (weight) is at most *k*.

Theorem

There is an algorithm that lists the Pareto-Optimal minimal adjustment sets for computing an unbiased estimator of the interventional mean of outcomes Y under interventions on X, of size at most k, in total time $O(4^k \cdot k \cdot (n+m))$.

Generating the Pareto-Optimal Frontier of Adjustment Sets

Fixed setting: DAG G; obtain unbiased estimator of the interventional mean of outcomes Y under interventions on X.

Ingredients:

- 1. Translate problem of finding adjustment sets in DAG G to finding *separators* in an undirected graph \mathcal{H} .
 - Validity constraints we want only adjustment sets $Z \subseteq R$.
 - Inclusion constraints we want only adjustment sets $I \subseteq Z$.

2. Translate critierion for (partially) ordering adjustment sets based on the variance they yield $(Z_1 \leq_G^{\sigma} Z_2)$ to a graphical criterion of separators.

Ingredient #1: From Adjustment Sets to Separators

- Let \mathcal{H} be an undirected graph, and $X, Y \subseteq V(\mathcal{H})$.
- ▶ $S \subseteq V(\mathcal{H})$ is an X, Y-separator of \mathcal{H} if in the graph $\mathcal{H}-S$ there is no path from X to Y.
- ► *S* is a *minimal X*, *Y*-separator if none of its proper subsets are.
- S is a minimum X, Y-separator of H if |S| ≤ |S'| for every X, Y-separator S'.



Graph \mathcal{H} , vertex sets X, $Y \subseteq V(\mathcal{H})$.

We denote by $\mathcal{S}_{X,Y}(\mathcal{H})$ the minimal X, Y-separators in \mathcal{H} .

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Ingredient #1: From Adjustment Sets to Separators

We denote by $\mathcal{A}_{X,Y}(I, R, G)$ all subsets $I \subseteq Z \subseteq R$ that are adjustment sets for computing an unbiased estimator of the interventional mean of outcomes Y under interventions X. We denote by $\mathcal{A}_{X,Y}^{\text{MIN}}(I, R, G)$ the minimal adjustment sets.

Theorem (based on van der Zander et al. 2019, Smucler et al. 2021)

There is an undirected graph ${\mathcal H}$ derived from G where :

1. $\mathcal{A}_{X,Y}(I, R, G) \neq \emptyset$ if and only if X and Y are non-adjacent in \mathcal{H} .

2. $S \in \mathcal{A}_{X,Y}^{MIN}(I, R, G)$ if and only if $S \in \mathcal{S}_{X,Y}(\mathcal{H})$.

Ingredient #2: Graphical Criterion for \leq_{G}^{σ}

- Smucler et al. 2021 presented a graphical criterion for the univariate case - single treatment variable X and single outcome variable Y.
- We generalize to the case where X and Y are variable-sets, and show an equivalent criteria.
- For an X, Y-separator S in \mathcal{H} , we denote by:

$$\mathcal{C}_X(\mathcal{H}-S) \stackrel{\text{def}}{=} \{ v \in V(\mathcal{H}) : v \text{ is reachable from } X \text{ in } \mathcal{H}-S \}.$$

$$\blacktriangleright C_Y(\mathcal{H}-S) \stackrel{\text{def}}{=} \{ v \in V(\mathcal{H}) : v \text{ is reachable from } Y \text{ in } \mathcal{H}-S \}.$$

Theorem

Let $S_1, S_2 \in \mathcal{A}_{X,Y}^{MIN}(I, R, G)$, then:

$$S_1 \leq_G^{\sigma} S_2$$
 if and only if $\mathcal{C}_Y(\mathcal{H} - S_1) \subseteq \mathcal{C}_Y(\mathcal{H} - S_2)$.

Graphical Criteria for Domination of Adjustment Sets

Theorem

Let $S_1, S_2 \in \mathcal{A}_{X,Y}^{MIN}(I, R, G)$, then:

 $S_1 \leq_{\mathcal{G}}^{\sigma} S_2$ if and only if $\mathcal{C}_{\mathbf{Y}}(\mathcal{H} - S_1) \subseteq \mathcal{C}_{\mathbf{Y}}(\mathcal{H} - S_2)$.

- Causal model G; R ⊆ V(G) are observed; treatment variables X and outcome variables Y.
- Let \mathcal{H} be the undirected graph derived from G where $\mathcal{A}_{X,Y}^{\text{MIN}}(I, R, G) = \mathcal{S}_{X,Y}(\mathcal{H})$ van der Zander et al., Smucler et al. 2019, 2021.
- ▶ Let $Z \in \mathcal{S}_{X,Y}(\mathcal{H})$.
- ▶ By applying the theorem, we have that Z ∈ A^{MIN}_{X,Y}(I, R, G) belongs to the Pareto frontier if and only if

$$\mathcal{C}_Y(\mathcal{H}-Z')\subseteq \mathcal{C}_Y(\mathcal{H}-Z)\Longrightarrow |Z|<|Z'|$$

for every $Z' \in \mathcal{A}_{X,Y}(I, R, G)$.

Pareto Frontier of Adjustment Sets \equiv *Important* Separators

Definition (Important Separator Marx 2011) Let $S \subseteq V(\mathcal{H})$. We say that S is an important X, Y-separator if $S \in \mathcal{S}_{X,Y}(\mathcal{H})$, and for any other $S' \in \mathcal{S}_{X,Y}(\mathcal{H})$ it holds that: $\mathcal{C}_{\mathbf{Y}}(\mathcal{H}{-}S') \subset \mathcal{C}_{\mathbf{Y}}(\mathcal{H}{-}S) \Longrightarrow |S'| > |S|$ Ingredients 1 + 2 $S \in \mathcal{A}_{X,Y}^{MIN}(I, R, G)$ is Pareto optimal $\Leftrightarrow S \in \mathcal{S}_{X,Y}(\mathcal{H})$ and important.

Pareto Frontier of Adjustment Sets \equiv *Important* Separators

 $S \in \mathcal{A}_{X,Y}^{\mathrm{MIN}}(I, R, G)$ is Pareto optimal $\Leftrightarrow S \in \mathcal{S}_{X,Y}(\mathcal{H})$ and important.

Theorem (Marx 2011)

There are at most 4^k important X, Y-separators of \mathcal{H} whose size is at most k, and they can all be generated in time $O(4^k \cdot k \cdot (n+m))$.

Pareto Frontier of Adjustment Sets \equiv Important Separators

 $S \in \mathcal{A}_{X,Y}^{\mathrm{MIN}}(I, R, G)$ is Pareto optimal $\Leftrightarrow S \in \mathcal{S}_{X,Y}(\mathcal{H})$ and important.

Theorem (Marx 2011)

There are at most 4^k important X, Y-separators of \mathcal{H} whose size is at most k, and they can all be generated in time $O(4^k \cdot k \cdot (n+m))$.

Theorem

The Pareto-Optimal minimal adjustment sets of size at most k, can be generated in total time $O(4^k \cdot k \cdot (n+m))$.

Is This Enough ?

- This approach works only if k is small.
- Variable weights bounded by constant c.
- Not really ranked: if k = 10, then adjustment sets of size 10 may be returned before those of size 3.

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Is This Enough ?

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Theorem

Let G be a causal DAG with weight function $w: V(G) \to \mathbb{N}_{\geq 1}$, and X, Y $\subseteq V(G)$ disjoint. There is an algorithm that enumerates all valid adjustment sets for (X, Y) in order of **non-decreasing total weight**, breaking ties by \leq_{G}^{σ} (i.e., proximity to Y). The delay is $O(Kn \cdot m^{1+o(1)})$, where K is the size of the largest set listed.

Conclusion & Outlook

- Our enumeration algorithms generate a comprehensive (pareto-optimal) set of valid, cost-constrained adjustment sets.
- These outputs have the potential to serve as training data for ML models that predict high-quality adjustment sets.
- This approach enables data-driven discovery of patterns connecting graphical structure to estimator variance.
 - By analyzing these patterns, we may gain new insights into how the position and composition of adjustment sets affect estimator variance.

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Thank you!

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